ORF 245 Fundamentals of Statistics
Chapter 1
Probability

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Sept 2014

Slides last edited on September 19, 2014
Course Info

Prereqs: Three semesters of Calculus

Textbook: *Mathematical Statistics and Data Analysis* by John A. Rice

Grading:
- Homework: 45%
- Midterm 1: 15%
- Midterm 2: 15%
- Final: 20%
- Participation: 5%

Homework:
- Will be due every week at 5pm on Friday.
- Turn homework in via ORF245 drop box in Sherrerd Hall.
- The lowest homework grade will be dropped.

Midterms:
Midterms will be in-class on Wednesday of the 5th and 10th weeks.

Lectures:
Reading material will be posted in advance of each lecture. You are expected to read the reading material before lecture.

Slides:
The slides will be posted online. But, they are not a replacement for the lecture. They are just my notes to remind me what to say. You must go to lecture to hear what I have to say.
Statistics is about extracting meaningful conclusions from noisy data. Here’s an example.

Is there a warming trend? If yes, what is the rate of warming? Eventually, we will answer this question.
In this class we will do some statistical computing.

We will also want to be able to plot data.

For these tasks, we will use a computer programming language called Matlab.

Here's the code that was used to make the plot on the previous page.

```matlab
load -ascii 'mat_output'
t=mat_output(:,1);
avg=mat_output(:,2);
plot(t,avg,'.b');
xlabel('Date');
ylabel('Avg Temp (degrees F)');
title('Average Daily Temperatures at McGuire AFB');
```

Before delving deeply into statistics, we need to have a good grasp of what we mean by “noisy” data.

So, we lay the groundwork with some probability...
Sample Spaces and Events

When considering experiments *to be performed*, we need to define the *set of possible outcomes*.

The set is called the *sample space* and is usually denoted by $\Omega$.

A sample space can be either finite or infinite.

**Examples:**

- Toss a coin: $\Omega = \{H, T\}$.
- Roll a die: $\Omega = \{1, 2, 3, 4, 5, 6\}$.
- Roll a pair of dice: $\Omega = \{(i, j) \mid 1 \leq i \leq 6, 1 \leq j \leq 6\}$.
- How long from now until the next major earthquake in California: $\Omega = \{t \mid t \geq 0\}$.

Subsets of $\Omega$ are called *events*. They are usually denoted by capital letters like $A$ or $B$.

**Example:**

- Roll a pair of dice and consider the event that the sum is 4:
  
  $$A = \{(1, 3), (2, 2), (3, 1)\} \subset \Omega.$$
And, Or, Not, etc.

Intersection: Event \( A \) and \( B \) both occur is written \( A \cap B \).
Union: Event \( A \) or \( B \) occurs is written \( A \cup B \).
Complement: Event \( A \) does not occur is written \( A^c \).
Empty Set: The set containing no elements is denoted \( \emptyset \).
Disjoint: Sets \( A \) and \( B \) are disjoint if \( A \cap B = \emptyset \).

Probability

*Probabilities* are numbers assigned to events. They must satisfy the following properties:

- \( P(\Omega) = 1 \).
- \( P(A) \geq 0 \) for all \( A \subset \Omega \).
- If \( A_1 \) and \( A_2 \) are disjoint, then \( P(A_1 \cup A_2) = P(A_1) + P(A_2) \).

It follows that

- \( P(A^c) = 1 - P(A) \).
- \( P(\emptyset) = 0 \).
- \( P(A) \leq P(B) \) whenever \( A \subset B \).
Examples:

1. Flip two pennies and let \( \Omega = \{(h, h), (h, t), (t, h), (t, t)\} \). Then,

\[
P(\text{one head and one tail}) = P(\{(h, t)\} \cup \{(t, h)\}) = P(\{(h, t)\}) + P(\{(t, h)\})
\]
\[
= \frac{1}{2} \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{2}
\]

2. Flip two pennies and let \( \Omega = \{\text{“two heads”}, \text{“one head”}, \text{“no heads”}\} \). Then,

\[
P(\text{one head and one tail}) = \text{hmmm... hard to say}
\]

3. Roll a pair of dice and let \( \Omega = \{(i, j) \mid 1 \leq i \leq 6, 1 \leq j \leq 6\} \). Then,

\[
P(\text{sum is 4}) = P(1, 3) + P(2, 2) + P(3, 1) = \frac{3}{36}
\]
**Conditional Probability**

**Definition.** The conditional probability of $A$ given that $B$ is known to have occurred is

$$P(A|B) = \frac{P(A \cap B)}{P(B)}.$$

**Law of Total Probability.** Let $B_1, B_2, \ldots, B_n$ be a disjoint collection of sets each having positive probability whose union is all of $\Omega$. Then,

$$P(A) = \sum_{i=1}^{n} P(A|B_i)P(B_i).$$

**Bayes' Rule.** If, in addition to the assumptions above, $P(A) > 0$, then

$$P(B_j|A) = \frac{P(A|B_j)P(B_j)}{\sum_{i=1}^{n} P(A|B_i)P(B_i)}.$$

If the disjoint collection consists of just two sets, $B$ and $B^c$, then the formula can be written more simply as

$$P(B|A) = \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|B^c)P(B^c)}.$$
Consider women of a certain given age and overall health status.

**Facts:**

- Probability that a woman has breast cancer \( = P(B) = 1\% \).
- Probability that a mammogram will give a positive result (indicating cancer is present) \((\text{event } A)\) for women who are known to have cancer \( = P(A|B) = 80\% \).
- A woman who is known not to have cancer will test positive \( 10\% \) of the time. That is, \( P(A|B^c) = 10\% \).

**Question:** If a woman tests positive for breast cancer, what is the probability that she actually has breast cancer? That is, what is \( P(B|A) \)?

Cancer doctors were asked this question. Most estimated the answer to be \( 75\% \).
Example of Bayes’ Rule

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Let’s compute:

\[
P(B|A) = \frac{(0.8)(0.01)}{(0.8)(0.01) + (0.1)(0.99)} = \frac{8}{8 + 99} \approx 7.5\%.
\]
Venn Diagram
Events $A$ and $B$ are independent means

$$P(A) = P(A|B)$$

$$P(B) = P(B|A)$$
Events $A$ and $B$ are independent means

\[ P(A) = P(A|B) = \frac{P(A \cap B)}{P(B)} \]

\[ P(B) = P(B|A) = \frac{P(A \cap B)}{P(A)} \]

\[ P(A \cap B) = P(A)P(B) \]
In a game of poker, what is the probability that a five-card hand will contain
(a) a straight (five cards in a numerical sequence not all from same suit),
(b) four of a kind (four cards of one value), and
(c) a full house (three cards of one value and two cards of another)?

(a) 
\[
\frac{10 \cdot (4^5 - 4)}{(52 \cdot 51 \cdot 50 \cdot 49 \cdot 48)/(5 \cdot 4 \cdot 3 \cdot 2 \cdot 1)} = 0.00392
\]

(b) 
\[
\frac{13 \cdot 48}{\binom{52}{5}} = 0.000240
\]

(c) 
\[
\frac{13 \cdot 12 \cdot \binom{4}{3} \binom{4}{2}}{\binom{52}{5}} = 0.001441
\]
Example E: Birthday Problem

In a classroom of \( n \) students, what’s the probability \( p_n \) that two (or more) students share the same birthday?

It’s easier to compute the probability that no two students share a birthday. Let’s look at the students one at a time. The first student can have any birthday he/she likes. The second student cannot share the first student’s birthday: 364 choices. The third student cannot share either of the first two birthdays: 363 choices. ... Etc. ...

The \( n \)-th student cannot share any of the previous \( n - 1 \) birthdays: \( 365 - n + 1 \) choices. Therefore, probability of no shared birthdays is

\[
\frac{365 \cdot 364 \cdots (365 - n + 1)}{365^n}
\]

and the probability of a shared birthday is

\[
p_n = 1 - \frac{365 \cdot 364 \cdots (365 - n + 1)}{365^n}.
\]

For \( n = 23 \), the answer is

\[
p_{23} = 0.507.
\]
Suppose you’re on a game show, and you’re given the choice of three doors: Behind one door is a car; behind the others, goats. You pick a door, say No. 1, and the host, who knows what’s behind the doors, opens another door, say No. 3, which has a goat. He then says to you, “Do you want to pick door No. 2?” Is it to your advantage to switch your choice?
A couple has two children.

(a) What is the probability that both are girls given that the oldest is a girl?

(b) What is the probability that both are girls given that one of them is a girl?

Let \( \Omega = \{(f, f), (f, m), (m, f), (m, m)\} \) (gender of older followed by gender of younger).

Let \( A = \text{“both are girls”} = \{(f, f)\} \).

Let \( B = \text{“oldest is a girl”} = \{(f, f), (f, m)\} \).

Let \( C = \text{“at least one is a girl”} = \{(f, f), (f, m), (m, f)\} \).

(a) \( P(A|B) = 1/2 \).

(b) \( P(A|C) = 1/3 \).
Show that if $A$ and $B$ are independent, then $A$ and $B^c$ are independent and so are $A^c$ and $B^c$.

Given: $P(A \cap B) = P(A)P(B)$.

Compute:

$$P(A \cap B^c) = P(A) - P(A \cap B)$$
$$= P(A) - P(A)P(B)$$
$$= P(A)(1 - P(B))$$
$$= P(A)P(B^c)$$

$$P(A^c \cap B^c) = P((A \cup B)^c)$$
$$= 1 - P(A \cup B)$$
$$= 1 - P(A) - P(B) + P(A \cap B)$$
$$= 1 - P(A) - P(B) + P(A)P(B)$$
$$= (1 - P(A))(1 - P(B))$$
$$= P(A^c)P(B^c)$$
Frequentist Approach – Coin Tossing

The graphs above illustrate the fraction of heads observed in coin tosses as a function of the number of tosses.

The top graph shows the results over 1000 tosses, demonstrating the convergence to a stable fraction of heads.

The bottom graphs zoom into the early stages of the experiment, highlighting fluctuations before reaching stability.

These diagrams are typical of what you would expect when performing a large number of trials, with the fraction of heads approaching the theoretical probability of a fair coin, which is 0.5.
ht = randi([0,1],[1,1000]);
St = cumsum(ht);
Xt = St./(1:1000);

figure(1);
plot(1:1000,Xt,'k-');
ylim([-0.02 1.02]);
xlabel('number of tosses');
ylabel('fraction of heads');